APPENDIX A

Detailed Reweighting Functions for Decomposition

In the text, we obtained the reweighting function conditional upon setting *all* individual attributes to 1940 levels simultaneously. In the context of this investigation, it is useful to partition the vector of individual characteristics, (X, t_x), into four subsets, $(Z, t_z; E, t_e; O, t_o; M, t_m)$. Z contains a vector of individual attributes including eight age categories, six hours categories, and indicator variables for marital status and the presence of a child at home. E represents a set of five dummy variables for educational attainment. The elements of O include eleven occupational and ten industrial categories. M denotes a set of indicator variables for the nine census regions and metropolitan residence. As before, t_j, denotes a binary random variable for the year of observation for each subset of characteristics.

Beginning with the observed 1950 earnings distribution, we adjust each set of characteristics to reflect those of observed 1940 distribution. We, first, reweight the distribution in personal characteristics (Z), then the distribution of educational attainment (E), then the distribution of occupations and industries (O), and finally the regional and metropolitan distribution of workers (M). At each step along the way, the new counterfactual distribution allows us to assess how changes in a subset of characteristics transformed the overall distribution of wages, all else held constant. The sequence of adjustments matters quantitatively because each subcomponent's contribution is measured in relation to the counterfactual distribution generated in the previous step. We selected this particular sequence with the idea that we are first adjusting for changes in the worker's basic characteristics and education level (primarily pre-market factors), and then for how workers are sorted once in the labor market (occupationally, industrially, and geographically).

As in equation (2), a counterfactual density of 1950 wages that is adjusted to reflect the 1940 distribution of Z characteristics, while maintaining the residential and occupational distribution of 1950, and maintaining the 1950 wage structure, can be written as:[79]

$$f(w; t_w = 50, t_z = 40, t_e = 50, t_o = 50, t_m = 50) = E_{Z,E,O,M}\left\{f_{50}(w \mid Z, E, O, M, t_w = 50) \times \psi_z(Z)\right\} \qquad \text{(A1)}$$

where the reweighting function is defined as

[79] This follows from rewriting $dF_X(x, t_x)$ in equation (2) using Bayes' theorem as a product of conditional densities: $dF_X(x, t_x) = dF_{Z,M,O}(z, m, o, t_z, t_m, t_o) = dF_{Z|M,O}(z \mid m, o, t_z, t_m, t_o) \times dF_{M|O}(m \mid o, t_m) \times dF_O(o \mid t_o)$.

$$\psi_z(z) = \frac{dF_{Z|E,O,M}(z \mid e,o,m,t_{z|e,o,m}=40)}{dF_{Z|E,O,M}(z \mid e,o,m,t_{z|e,o,m}=50)} = \frac{P(t_{z|e,o,m}=40 \mid z,e,o,m)}{P(t_{z|e,o,m}=50 \mid z,e,o,m)} \cdot \frac{P(t_{z|e,o,m}=50 \mid e,o,m)}{P(t_{z|e,o,m}=40 \mid e,o,m)}. \quad \text{(A2)}$$

Similarly, the counterfactual density of wages, that would have prevailed if individual attributes *and* education were distributed as in 1940, while maintaining the occupational, industrial, geographic and wage distribution of 1950, can be written as:

$$f(w;t_w=50,t_z=40,t_e=40,t_o=50,t_m=50) = E_{Z,E,O,M} f_{50}(w \mid Z,E,O,M,t_w=50) \times \psi_z(Z) \times \psi_e(E) \quad \text{(A3)}$$

where the reweighting function is defined

$$\psi_e(e) = \frac{dF_{E|O,M}(e \mid o,m,t_{e|o,m}=40)}{dF_{E|O,M}(e \mid o,m,t_{e|o,m}=50)} = \frac{P(t_{e|o,m}=40 \mid e,o,m)}{P(t_{e|o,m}=50 \mid e,o,m)} \cdot \frac{P(t_{e|o,m}=50 \mid o,m)}{P(t_{e|o,m}=40 \mid o,m)}. \quad \text{(A4)}$$

Finally, the counterfactual occupational and geographic distributions are

$$f(w;t_w=50,t_z=40,t_e=40,t_o=40.t_m=50) \quad \text{(A5)}$$
$$= E_{Z,E,M,O} f_{50}(w \mid Z,E,M,O,t_w=50) \times \psi_z(Z) \times \psi_e(e) \times \psi_o(O) \text{ and}$$
$$f(w;t_w=50,t_z=40,t_e=40,t_o=40.t_m=40) \quad \text{(A6)}$$
$$= E_{Z,E,M,O} f_{50}(w \mid Z,E,M,O,t_w=50) \times \psi_z(Z) \times \psi_e(e) \times \psi_o(O) \times \psi_m(m), \text{ and}$$

can be obtained by applying the reweighting functions,

$$\psi_o(o) = \frac{dF_O(o \mid m,t_{o|m}=40)}{dF_O(o \mid m,t_{o|m}=50)} = \frac{P(t_o=40 \mid o,m)}{P(t_o=50 \mid o.m)} \cdot \frac{P(t_o=50 \mid m)}{P(t_o=40 \mid m)} \text{ and} \quad \text{(A7)}$$

$$\psi_m(m) = \frac{dF_M(m \mid t_m=40)}{dF_M(m \mid t_m=50)} = \frac{P(t_m=40 \mid m)}{P(t_m=50 \mid m)} \cdot \frac{P(t_m=50)}{P(t_m=40)}. \quad \text{(A8)}$$

Following the discussion above, the analysis of the absolute changes in the density of wages between 1940 and 1950 is based upon the sequential decomposition within each racial group described in the text.

APPENDIX B

Data and Sample Description for the 1940 and 1950 IPUMS

The 1940 and 1950 data samples are drawn from the IPUMS. We exclude a number of workers to maintain consistent and comparable samples for both periods. The samples include workers between the ages of 18 and 64, who were not in school, in the Armed Forces, or self-employed (because non-wage income is not reported in 1940); who were not farmers, farm managers, or farm workers; who did not reside in institutional group quarters; and who worked more than four weeks in the previous calendar year.

For analysis, the age categories group individuals as follows: 18 to 25, 25 to 29, 30 to 34, 35 to 39, 40 to 44, 45 to 49, 50 to 54, and 55 to 64 years old. The hours categories are: no hours in the previous week (but with reported income for the previous year), 1 to 20 hours, 21 to 30, 31 to 40, 41 to 50, and more than 50 hours. Dummies for educational attainment pertain to: those with less than five years, five to eight, nine to 11, exactly 12 years and more than 12 years of schooling. Occupational groups are: professionals, operatives, laborers, household workers, service workers, managers, sales persons, clerical workers, crafts persons, or newly employed (without an occupation in the previous year). Industrial groups are: agriculture/forestry/fishing/mining/construction, durable manufacturing, nondurable manufacturing, transportation/communication/utilities, trade, finance/insurance/real estate, business and personal services, professional services, entertainment/recreation services, and public administration. The omitted category consists of unmarried women, ages 35 to 39, with 9 to 11 years of schooling, who worked as non-durable manufacturing operatives for 31 to 40 hours in the week before the census, and lived in the East North Central census division.

APPENDIX C

Log Weekly Wage Regressions by Race, 1940 and 1950

	White women		Black women	
	1940	1950	1940	1950
<=4 years	-0.177	-0.104	-0.193	-0.147
	(9.23)	(3.13)	(8.13)	(3.75)
5-8 years	-0.071	-0.067	-0.065	-0.058
	(12.48)	(5.50)	(3.87)	(2.17)
12 years	0.049	0.075	0.043	0.045
	(7.98)	(8.41)	(1.67)	(1.99)
13 or more years	0.251	0.2	0.22	0.179
	(22.43)	(15.96)	(6.52)	(4.17)
Married, spouse present	0.054	0.035	0.004	-0.033
	(9.39)	(5.48)	(0.30)	(2.53)
Child	-0.079	-0.05	-0.024	-0.007
	(10.72)	(8.57)	(2.61)	(0.46)
No hours worked	0.006	-0.099	-0.056	-0.088
	(0.24)	(4.54)	(1.38)	(1.39)
1-20 hours	-0.229	-0.456	-0.112	-0.407
	(21.05)	(28.29)	(5.50)	(10.66)
21-30 hours	-0.095	-0.183	-0.075	-0.167
	(11.51)	(11.16)	(5.16)	(7.14)
41-50 hours	-0.023	-0.032	0.042	-0.067
	(4.56)	(3.85)	(4.14)	(3.02)
50 hours	-0.074	-0.107	0.103	0.001
	(5.91)	(7.31)	(6.49)	(0.03)
<=25 years old	-0.32	-0.17	-0.194	-0.098
	(26.43)	(14.02)	(9.05)	(3.36)
25-29 years old	-0.131	-0.034	-0.1	-0.035
	(16.83)	(3.17)	(6.42)	(1.20)
30-34 years old	-0.035	-0.022	-0.025	0.004
	(4.85)	(2.05)	(1.79)	(0.18)
40-45 years old	0.027	0.026	0.006	0.024
	(3.41)	(2.29)	(0.33)	(0.93)
45-49 years old	0.026	0.026	0.044	0.005
	(3.61)	(2.00)	(2.00)	(0.14)
50-54 years old	0.028	0.028	0.047	-0.05
	(3.23)	(1.93)	(2.67)	(1.56)
55-65 years old	-0.029	-0.029	0.012	-0.092
	(3.16)	(1.61)	(0.50)	(3.52)
New England	0.013	-0.063	0.161	-0.063
	(0.58)	(2.38)	(11.22)	(0.96)
Middle Atlantic	0.052	0.018	0.109	0.015
	(1.22)	(0.36)	(2.53)	(0.26)
West North Central	-0.136	-0.075	-0.161	-0.243
	(4.62)	(3.51)	(5.64)	(6.17)
South Atlantic	-0.025	-0.032	-0.313	-0.333
	(0.96)	(1.34)	(3.63)	(4.45)

East South Central	-0.151	-0.154	-0.546	-0.551
	(7.97)	(6.81)	(7.83)	(11.94)
West South Central	-0.17	-0.119	-0.435	-0.39
	(9.12)	(4.73)	(14.23)	(8.81)
Mountain	-0.001	-0.011	0.09	0.145
	(0.03)	(0.29)	(0.91)	(1.13)
Pacific	0.112	0.082	0.2	0.133
	(3.82)	(3.46)	(10.72)	(3.44)
Lives in city	0.234	0.158	0.36	0.261
	(16.42)	(11.87)	(9.98)	(8.64)
Professionals	0.402	0.236	0.585	0.561
	(34.15)	(13.95)	(9.51)	(7.03)
Clerical	0.157	0.026	0.196	0.025
	(17.14)	(2.20)	(3.40)	(0.48)
Craft	0.155	0.083	-0.157	0.029
	(11.05)	(4.62)	(2.06)	(0.41)
Laborer	-0.036	0.005	0.056	-0.018
	(1.97)	(0.15)	(0.97)	(0.29)
HH Service	-0.508	-0.47	-0.315	-0.25
	(16.78)	(15.32)	(12.47)	(8.71)
Manager	0.244	0.141	-0.169	-0.371
	(7.71)	(5.95)	(1.15)	(2.94)
Sales	-0.015	-0.134	-0.017	-0.158
	(1.35)	(9.13)	(0.16)	(1.86)
Unemployed	-0.161	-0.112	-0.124	-0.152
	(9.71)	(3.38)	(2.25)	(2.00)
Not in labor force	-0.156	-0.111	-0.211	-0.17
	(5.68)	(1.93)	(1.75)	(0.75)
Service, not HH	-0.155	-0.194	-0.149	-0.034
	(14.24)	(14.14)	(4.84)	(1.31)
Ag, forest, fish; mining, construct.	0.011	-0.008	0.232	-0.195
	(0.50)	(0.19)	(4.58)	(1.55)
Durable manuf.	0.071	0.047	-0.027	-0.042
	(3.33)	(2.36)	(0.29)	(0.78)
Trans, comm.., utility	0.07	0.073	0.009	0.059
	(4.81)	(5.92)	(0.08)	(0.69)
Trade	-0.094	-0.094	-0.225	-0.171
	(7.78)	(7.31)	(2.64)	(2.93)
Finance, ins., real est.	-0.005	-0.089	-0.264	-0.155
	(0.46)	(7.52)	(2.53)	(2.54)
Business and pers. services	-0.159	-0.156	-0.297	-0.197
	(8.90)	(9.33)	(3.22)	(3.74)
Ent. and rec. services	-0.025	-0.142	-0.081	-0.142
	(0.47)	(3.21)	(0.72)	(1.44)
Prof. services	-0.049	-0.084	-0.169	-0.124
	(2.92)	(6.85)	(2.10)	(1.80)
Pub. Administration	0.106	0.045	-0.049	0.17
	(5.06)	(2.21)	(0.32)	(2.07)
No industry, but wages	-0.057	-0.098	-0.186	-0.222
	(3.41)	(2.43)	(1.68)	(0.93)

Constant	3.228	3.587	2.981	3.435
	(148.31)	(154.02)	(30.92)	(48.86)
Observations	86166	39597	12131	5413
R-squared	0.42	0.25	0.49	0.42

Robust t statistics are in brackets. Omitted categories are women ages 35-39 with 9-11 years of schooling who were working 31-40 hours per week as operatives in nondurable manufacturing and lived in the East North Central Division. Source: Authors calculations from the 1940 and 1950 IPUMS samples (Ruggles et al. 2004).

APPENDIX D

Descriptive Statistics for the March *CPS* Samples, 1963-2001

A. Sample averages by year, for all women over 16 to 65

	IPUMS					March *CPS*				
	1940	1950	1960	1965	1970	1975	1980	1985	1990	2000
In the labor-force	0.242	0.306	0.362	0.435	0.489	0.527	0.590	0.632	0.676	0.710
Hours worked	41.36	38.39	35.87	35.84	34.69	34.16	34.68	35.37	36.92	37.55
Weeks worked	39.2	37.1	36.0	38. 3	39.2	40.3	40.6	42.1	44.5	47.0
Currently married	0.658	0.710	0.715	0.720	0.700	0.665	0.626	0.600	0.638	0.668
Never married	0.252	0.167	0.167	0.155	0.170	0.184	0.210	0.223	0.168	0.109
16 or more years of schooling	0.036	0.052	0.053	0.071	0.081	0.103	0.129	0.154	0.199	0.262
Age	36.3	37.6	38.5	38.2	37.7	37.3	37.0	37.1	39.4	45.1

B. Sample averages by year, for all women over 16 to 45

	IPUMS					March *CPS*				
	1940	1950	1960	1965	1970	1975	1980	1985	1990	2000
In the labor-force	0.237	0.301	0.356	0.435	0.500	0.562	0.646	0.694	0.743	0.775
Hours worked	41.3	38.3	35.7	35.1	34.0	33.7	34.6	35.2	37.2	37.6
Weeks worked	39.2	36.9	35.9	36.2	37.1	38.7	39.5	41.2	44.3	46.9
Currently married	0.642	0.697	0.699	0.727	0.697	0.646	0.593	0.563	0.618	0.673
Never married	0.270	0.182	0.185	0.198	0.226	0.250	0.283	0.299	0.225	0.146
16 or more years of schooling	0.035	0.051	0.052	0.071	0.084	0.116	0.145	0.172	0.224	0.282
Age	35.8	37.2	37.9	30.0	29.2	28.7	28.7	29.4	32.0	37.6

Sample: Women not in the military or inmates. Sources: March *CPS* 1964-2001 and 1940-1960 PUMS (Ruggles et al 2004).

APPENDIX E

Fraction of Women with Early Access for Selected Birth Cohorts

June CPS 1977-1995

Year of birth	Fraction of ever-married women with early access (Number of observations in June *CPS*)
1935	0.000
	(2,629)
1940	0.051
	(6,255)
1945	0.066
	(7,002)
1950	0.132
	(6,729)
1951	0.429
	(5,778)
1955	0.977
	(2,908)
1960	1.000
	(77)

March CPS 1965-1995

	Year of observation						
Year of birth	1965	1970	1975	1980	1985	1990	1995
1935	0.000	0.000	0.000				
	(411)	(816)	(671)				
1940	0.065	0.061	0.077	0.040			
	(450)	(821)	(762)	(960)			
1945	0.066	0.080	0.079	0.071	0.080		
	(542)	(1,014)	(877)	(1,236)	(1,006)		
1950	0.101	0.149	0.149	0.140	0.144	0.101	
	(688)	(1,177)	(1,046)	(1,438)	(1,301)	(1,209)	
1951		0.252	0.434	0.431	0.435	0.432	0.439
		(1,188)	(1,045)	(1,510)	(1,443)	(1,269)	(1,080)
1955		0.300	0.977	0.978	0.979	0.980	0.300
		(1,463)	(1,167)	(1,655)	(1,354)	(1,351)	(1,183)
1960			1.000	1.000	1.000	1.000	1.000
			(1,354)	(1,652)	(1,472)	(1,416)	(1,221)

Notes: The table cells contain the fraction of women who would have had access under the definition in equation (2) by year of birth and year of observation in the March *CPS*. The number of observations is in brackets. Samples correspond to those in Table 2 and Table 4 respectively.

APPENDIX F

The Aggregate and Long-term Effects of Early Access to the Pill, Counterfactual Estimates 1960-1990

		1940	1950	1960	1965	1970	1980	1990
A. Observed participation rates								
	Participation of 16-65	0.242	0.306	0.362	0.435	0.489	0.590	0.676
	Participation of 16-30 (PT1)			0.356	0.392	0.465	0.617	0.717
	Participation of 16-45 (PT2)			0.356	0.414	0.473	0.625	0.734
B. No access counterfactual								
(1)	No access ages 16-30 (NA1)				0.392	0.463	0.599	0.682
	No access ages 16-45 (NA2)				0.414	0.472	0.613	0.716
					1960-65	1965-70	1970-80	1980-90
(2)	Percentage points attributed to access							
	Women ages 16-30	(PT1-NA1)			<0.001	0.002	0.018	0.035
	Women ages 16-45	(PT2-NA2)			<0.001	0.001	0.011	0.018
(3)	Percent increase from t-1 to t attributed to access							
	Women ages 16-30				0.007	0.022	0.121	0.350
	Women ages 16-45				0.003	0.016	0.074	0.162

The no access counterfactual in Panel B simulates the state of the world if no woman, from 1960 to the present, had gained legal access to the pill before her 21st birthday. Using the estimates obtained for the model in equation (3), I predict individual participation rates and average over the particular year and age group to obtain the estimates in line 1. Line 2 presents the difference in percentage points between the observed and predicted participation rates. Line 3 transforms these numbers into percentage changes dividing (PT1-NA1) by (NA1(t)-NA1(t-L)) where t is the year and L denotes either a 5 or 10 year date difference. Sources: Estimated effects based on 1964-2001 March *CPS*; observed participation rates from 1940-1960 PUMS (Ruggles et al. 2004) and 1965-1990 March *CPS*.

APPENDIX G

Sample Restrictions and Data Construction for 1960-200 IPUMS

Weeks worked:

Prior to 1980, intervals rather than actual values were reported for weeks worked. Rather than assigning the mid-points of the intervals as values for observations prior to 1980, I use the median number of weeks worked in the same interval in the 1980 census. The results of the two methods are compared below.

	Weeks worked		
1960-70 Intervals	1960, 1970 Values	Midpoint	Median 1980
1-13	1	7	8
14-26	2	20	20
27-39	3	33	32
40-47	4	43.5	42
48-49	5	48.5	48
50-52	6	51	52

Deflator:

I use the CPI numbers recommended by IPUMS and report all figures in 1990 dollars.

Earnings Top Codes:

The standard methodology to adjust for top-coding is to multiply the top-code by some standard factor, for instance 1.5 (Acemoglu et al. 2004) or 1.45 (Autor and Katz 1999; Juhn, Murphy, and Pierce 1993; Goldin and Margo 1992). In order to obtain better measures of within state inequality, I improve upon this methodology in two ways.

First, there is considerable information to be gained by exploiting the fact that the top-code is not indexed to inflation. For instance, the nominal top-code in 1970 is $50,000 and the top-code for 1960 in 1970 dollars is $31,415. Under the assumption that the mean earnings in the interval, $31,415 to $50,000, remained unchanged over the 1960s, we can infer the distribution of earnings for individuals above the 1960 top-code in the 1960 census. I do this for individuals in 1980, 1990 and 2000 as well. I plot the real value of the top-code in 1990 dollars in Figure A.

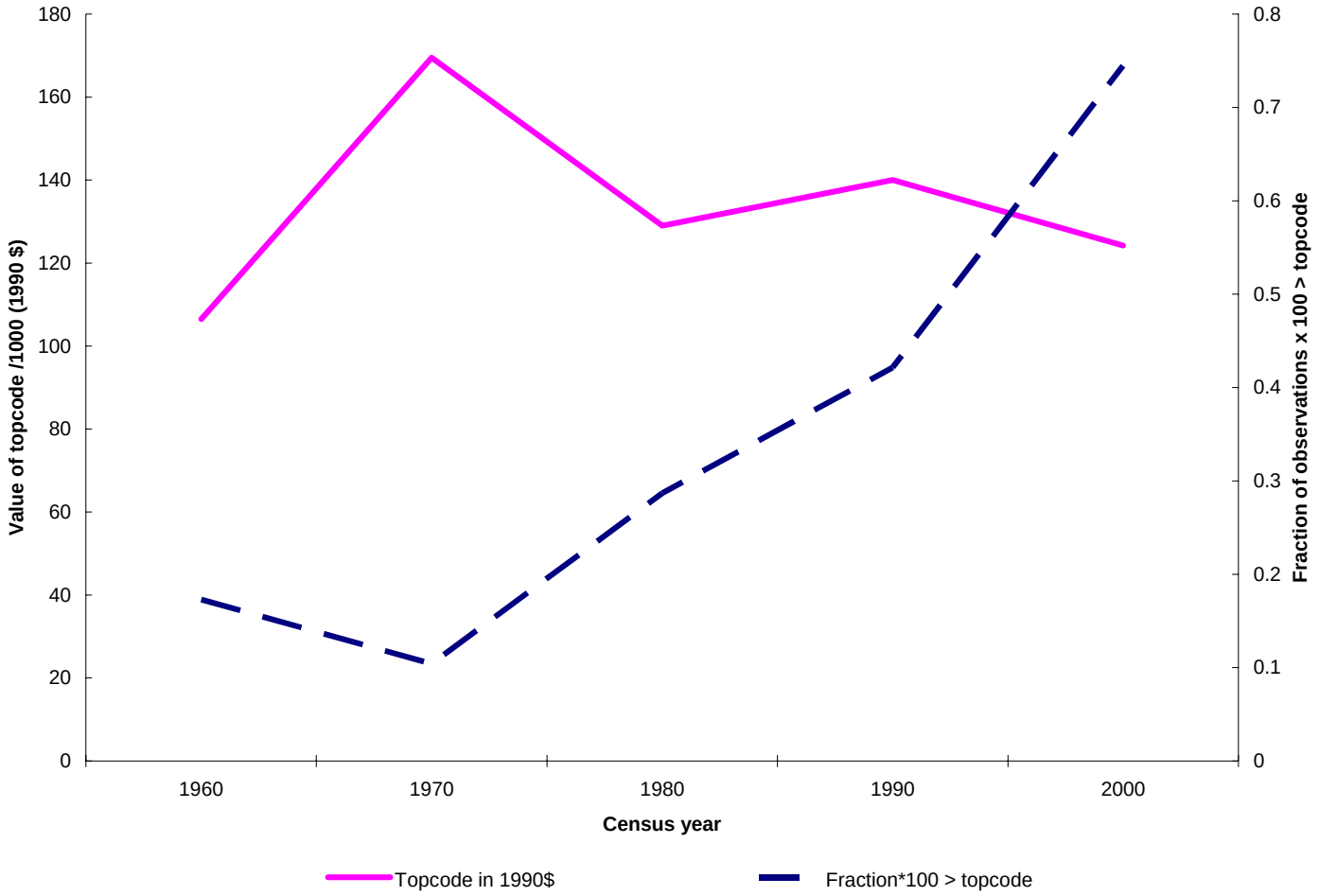

Figure A. The Real Value of the Top Code for Earnings Data in the Census, 1960 to 2000

Second, I make use of cross-state variation in the mean earnings of individuals above the census top-code. Individuals with top-coded earnings receive the *state* mean rather than the aggregate mean.

I use the following procedure. Let l denote the census year's top-code for any year, y = 1960, 1980, 1990, 2000, and u denote the 1970 top-code. All of the values have been adjusted to 1990 dollars as in the analysis. I multiply all observations with the 1970 top-code by 1.5. Then—by state—I replace all observations equal to l with the mean value taken over the observations in the interval $[l, u]$.

Sample Restrictions:

In each census year, the sample is restricted to individuals born in the U.S. who are ages 16 to 65, who reported positive earnings, and who were not residing in group quarters. Outliers are excluded by additionally limiting the sample to individuals who either earned at least $1 per hour and no more than $100 per hour (see DiNardo, Fortin and Lemieux 1996, Fortin and Lemieux 1998) or an equivalent amount in weekly terms, $40 to $4,000 in earnings on an average annual weekly basis.

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